"All in the Family: Do Health Utility Patterns Cluster in Belgian Households?"

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Abstract

Evidence has shown that an individual's health and well-being can have quantifiable spillover effects on health-related quality of life (HRQoL) of caregivers and/or family members. Many studies have focused on individual-level data to explore factors influencing HRQoL, but few have evaluated within- and between household variations in HRQoL. Our work aims to explore clustering patterns of health utility among and within households in Belgium and identify individual and household-level determinants of HROoL that can influence these clustering patterns. A Belgian national survey collected data on 1,336 individuals in 340 households in 2010-11 to study social contact behavior of households with at least one child under 13 years old. The survey collected participants' individual and household characteristics and EQ-5D-3L. A linear mixed-effects model was used to identify determinants influencing individual EQ-5D utility scores, while accounting for clustering within households. A total of 1,212 individuals over 307 households were included in this analysis. The median household size was 4, with a range of 2 to 7, and there were 153 households with at least one member with a health utility score less than 1. The average intraclass correlation coefficient (ICC) for the final model from 10 imputed datasets was 11.7% (range: 11.1%-2.0%), indicating modest heterogeneity in health utility scores across households. The individual-level predictors found to have a significant association with lower individual health utility scores include being sick with a severe illness (b = -0.06, p < 0.001), having dealt with a family member with a severe illness (b = -0.03, p < 0.05), and experiencing an abnormal day at time of response (b = -0.02, p < 0.05). Participants whose BMI was classified as 'obese' (b = -0.07, p < 0.001), 'overweight' (b = -0.03, p < 0.05), and 'underweight' (b = -0.04, p < 0.05) showed significant associations with lower health utility scores. Participants with 'other' occupational status (i.e., housewife, unemployed, or job-seeking; b = -0.066, p < 0.001) were found to have a significant negative association. Household clusters have a non-negligible influence on health utility scores that requires further investigation. Lower health utility scores among individuals are greatly associated with having or being in close proximity to severe diseases.

Introduction

Health is defined by a person's overall physical, mental, and social well-being and is influenced by a wide range of biological, societal, and environmental factors (1). An individual's health has been shown to have direct and indirect influences on the health and well-being of other people within close social relationships (in a community or household, etc) (2). These are apparent in most common infectious diseases with physical manifestations. Several studies have directly measured the rate of infectious disease transmission in households in Belgium (3,4); however, research has shown that subtle spillover effects, that is, the impact that an individual's health condition has on the health and well-being of others (such as family members and caregivers) (5,6) , also occur with non-communicable diseases (NCDs) and mental health conditions (7). Furthermore, lifestyle habits that have direct influence on health and overall quality of life are also shared among household members (8,9) and can have downstream implications on society and overall public health.

Health-related quality of life (HRQoL), including health utility scores, of individuals or patients suffering from a variety of diseases have been widely documented and its use to measure the commutation of health due to NCDs has been ubiquitous. There is a breadth of literature on the shared burden of disease within close contact relationships and inclusion of spillover effects in economic evaluations and healthcare policy-making is increasing (10). Chronic diseases, such as diabetes, have been shown to impact parents' stress and mental well-being due to worries about the child's health (11,12) and the involvement in disease maintenance (13). Radicke et al (14) found that parents' mental illness is associated with lower HRQoL for both parents and their children. Alternatively, Lawson et al (15) suggests that both positive and negative moods, due to mother's experience at work, can consequently affect the mood of their children.

Informal care is the care provided by family to patients with chronic or lifelong diseases. According to a report from the Belgian Health Care Knowledge Centre, 17% of the Belgian population older than 50 years old are informal caregivers in 2018 (16). A study examining the demographics of informal caregivers in Belgium (17) observed higher levels of psychological distress with higher care intensity. There is currently a greater focus on informal caregiving for the elderly population due to aging populations globally (17). Informal caregiving, however, also pertains to adolescents who are expected to care for younger siblings or grandparents (19). Much of the current literature centers on parent-child, parent-elderly, and patient-caregiver relationships. Our work, which aims to explore clustering patterns of HRQoL among and within households with young children in the general Belgian population, is fundamental in filling research gaps on family spillover effects. We aim to identify individual and household-level determinants of HRQoL that can influence these clustering patterns, with consideration of household members' exposure to illness in the family or with someone being cared for.

Methods

Household Survey

A survey of Belgian households was conducted alongside a general population survey to study social contact behaviours in the Flemish population between November 2010 to April 2011 (3,4,20). Survey recruitment and dissemination are outlined in Goeyvaerts et al (3). In addition to contact diaries, this survey consisted of a general background questionnaire and a generic HRQoL questionnaire (EQ-5D-3L) with a Visual Analogue Scale (VAS) completed by all members of recruited households, which included at least one child aged 12 years or less. Two types of surveys were used - a normal adult survey and a similar survey adapted for children aged 0 to 12 years old with less formal language and completed by the parent or guardian in the household. Teenagers aged 13 to 18 years were asked to complete the survey on their own. A total of 1,336 individuals from 340 households participated in the survey.

Data

<u>Table 1</u> lists the subset of 23 survey questions that were selected for inclusion in this analysis based on prior literature and relevance to the research topic. Variables that were recategorized or created based on the original data are indicated in the table.

Health Related Quality of Life

The survey responses to the five EQ-5D dimensions were used to calculate the primary outcome of interest, an individual health utility score using the EQ-5D-3L value set for Belgium (21)('eq5d' R package version 0.15.3). Individuals in the survey were included based on

non-missing responses to each of the five EQ-5D dimensions. Participants' VAS health score was also ascertained in the survey.

Data Transformation and Imputation

The percentage of missing responses for the final included individuals based on EQ-5D are outlined in <u>Table 1</u>. A binary variable 'child aged under 13 - yes or no' was created due to the inherent age-specific recording of survey responses. Missing values for 'animal ownership' were manually re-coded with responses ascertained from responses by other household members, given that all other non-missing responses within the household are consistent. These values are assumed to be consistent among all members of the household as individuals are included in the survey if they live more than 50% of the time in a household.

Household-level variables for child care, severe disease, and smoking were also created. A household is determined to have child care if at least one child attends day care or school. Household-level exposure to severe disease was determined if any individual household member responded 'yes' to having dealt with severe illness themselves, with a family member, or with someone in their care. If any individual household member reported being a 'current smoker', the household is determined to have at least 1 smoker, while other households are identified as having no smokers. Household-level educational attainment and occupation were ascertained for the original data (prior to imputation) from adult members of the household. Household education was defined as having at least one parent with a university degree or both parents having less than a university-level education. The binary household occupation variable was defined as having both parents working full- or part-time or having at least one parent not working full- or part-time.

Missing values were handled with multiple imputation using the R package 'mice' (version 3.16.0) unless otherwise denoted. Given the wide range of proportions of missing values (1% to 33%), a total of 10 imputation iterations were implemented with 75 burn-in iterations. Variables with missing values were assessed for multicollinearity against remaining possible predictor variables for imputation. Height and weight were regressed against age, gender, height and weight to account for changes in growth rates for children and adults. The remaining variables with missing values were regressed against all other original variables except variables deemed multicollinear in initial imputation variable specification and identifier variables. Binary

variables were imputed using logistic regression and remaining variables were imputed using the predictive mean matching method due to the nature of the variable or sparseness (22).

There were 10 new variables created after imputation that were further considered in our final model. The variable Body Mass Index (BMI) was calculated as BMI = weight in kg/(length in m)² and categorized into standard categories: underweight, normal weight, overweight, and obese, based on World Health Organization (WHO) guidelines. BMI categorization for participants under 18 years were determined using standard deviation scores and growth percentiles calculated using the 'childsds' package in R (version 0.8.0) referencing WHO standards or Belgian parameters, if applicable. New maternal education, adult education, and combined education variables were created to ascertain individual education levels from 'child school' and 'participant education' responses. The education and occupation variables were further recategorized to facilitate interpretation of results and to ensure adequate sample sizes in each category.

Statistical Analysis

Descriptive statistics for all potential determinants, using the original (hereinafter referred to as the baseline) and imputed datasets, are outlined in <u>Table 1</u>. The association between baseline averages of household health utility scores and household educational attainment and occupation status were assessed using linear models. Correlations between potential determinants were assessed using contingency tables, Spearman rank correlations, and chi-square tests. The associations between health utility scores and all potential determinants were examined using linear mixed effects models accounting for household-level clustering. For variables in which survey responses vary by age group, models were adjusted to include an interaction with a binary age group identifier and compared using AIC. Models with original and newly created variables were further manually compared using AIC and the variable with the lowest average AIC for all 10 imputation iterations was retained.

A fully specified model was identified after selecting between duplicated or correlated variables. Backward model selection was performed on each individual imputed data set using the fully specified model. For each imputed data set, the variables in the selected model were tallied and included in the model based on how many times a variable appeared in the results. The selected model was further assessed for multicollinearity and fit using a linear mixed-effect

model with household clusters as a random effect (23). Pooled regression coefficients for the final model were obtained using Rubin's rules.

Results

Baseline Sample Characteristics

Individuals in the survey were subsequently "filtered" to ensure that remaining households consisted of 2 or more individuals, with at least one child under 13 years old. A total of 1,212 individuals in 307 households were included in this analysis. The average age of participants included in the analysis is 23.5 (17.5) [0, 60] with 43% categorized as 'children' based on survey guidelines and 51% identifying as female and 49% identifying as male. There were 101 individuals who reported experiencing severe disease, 482 reported having contact with a family member with severe disease, and 63 reported caregiving for someone with severe disease. The average health utility score for the study sample was 0.94 [0.24, 1] and the average VAS health score for the study sample at baseline was 87.93 [7, 100].

Associations between Health Utility Score and Household Characteristics

Household sizes in our sample range from 2 to 7 members, with 92% of total households identifying as two-parent households and 8% identifying as single-parent households. There were 237 participants among 153 households with a health utility score less than 1. Out of the 153 households, there were 62 households with more than 1 member with health utility score less than 1 and 6 households in which all members' health utility score was less than 1. For households with at least 1 member with health utility scores below 1, the average health utility score is 0.88 [0.24, 1] and the average VAS health score is 85.6 [7, 100]. Figure 1 shows a distinct split in the distribution of average household health utility scores based on household composition, with households where more than 50% of members have health utility scores below 1 exhibiting lower average health utility scores.

Table 1 - Descriptive Statistics of Original Survey Results After Accounting for Households with Non-Missing Responses to theFive EQ-5D Dimensions

Variable	Variable Description	Levels	N (Baseline)	Percent Missing (out of 1212)
		Adult, himself	606	
Туре	Who filled the diary in?	Teenager or older child himself	81	0%
		Parent on behalf of child	525	
		Mean	23.5	
4	A so of porson whom the diamy refers to	Median	17.5	00/
Age	Age of person whom the diary refers to	Min	0	0%
		Max	60	
Age Group		Child (<= 12 years old)	517	00/
Category	Age group (adult or child)	Adult (> 12 years old)	695	0%
Candan	Conden of the norman where the diamy reference	Female	615	0%
Gender	Gender of the person whom the diary refers to	Male	597	0%
		Mean	1.52	
II ai alut	The bright of the nonticipant (in m)	Median	1.62	2.40/
Height	The height of the participant (in m)	Min	0.47	34%
		Max	2.01	
		Mean	52	
W 7-:-1-4	The main late of the manufacture of (in 1-2)	Median 55	250/	
Weight	The weight of the participant (in kg)	Min	6	35%
		Max	133	
		Mean	4.1	
Household Size	Household size including participant	Median	4	0%
		Min	2	

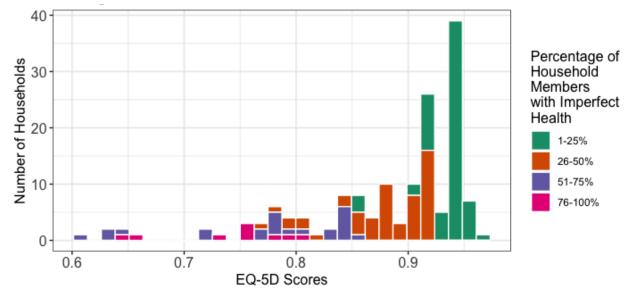
		Max	7	
Severe Disease -	Have you been in contact with a severe disease: yourself?	Yes	101	3%
Self	Have you been in contact with a severe disease. yoursen?	No	1078	570
Severe Disease -	Have you been in contact with a severe disease: family member?	Yes	482	4%
Family	Have you been in contact with a severe disease. failing member?	No	687	470
Severe Disease -	Have very have in contact with a second discount company were	Not applicable/child respondent	525	
Caregiving	Have you been in contact with a severe disease: someone you were taking care of?	Yes	63	7%
Caregiving		No	538	
Household	What kind of household are you living in?	Single-parent family	67	0%
Parental Type	what kind of nousehold are you nying in?	Two-parent family	1145	070
Uncommon Day	Was this an uncommon day?	Yes	361	1%
Uncommon Day	was uns an uncommon day?	No	844	1 70
Animal	Does your household own at least one living animal?	Yes	840	0%
Ownership	Does your household own at least one hving annual?	No	372	070
		Not applicable/adult respondent	687	
Child Care	Does the child attend daycare or school?	Yes	515	0%
		No	9	
		Not applicable/adult respondent	690	
		Yes, pre-school	129	
Child School	[<12y] Is the child going to school?	Yes, primary school	282	0%
		Yes, not pre- or primary school	56	
		No	53	
		No formal schooling	5	
	Educational level of the person filling in the diary (only if diary type	Primary school	52	
Education	== 1 or 3)	Secondary school (lower)	73	1%
	For children < 13 years : maternal education	Upper secondary school (upper)	261	
		Secondary school (unspecified)	100	

		University degree (lower)	462	
		University degree (higher)	247	
		University degree (unspecified)	0	
		Vocational education (FI)	0	
		Working full- or part-time	537	
		Retired	0	
		At home (housewife)	22	20/
Occupation	Occupational status of the person filling in the diary	Currently unemployed/job seeking	13	3%
		In full time or further education	550	
		Other	50	
		Not applicable/child respondent	525	
		Smoker	110	
Smoking	Individual smoking status	Ex-smoker	107	0%
		Non-smoker	465	
		Unknown	5	
		I have no problems in walking about	1193	
EQ-5D - Mobility	The health state of the participant: regarding mobility	I have some problems in walking about	18	0%
		I am confined to bed	1	
		I have no problems with self-care ^a	1181	
EQ-5D - Self-Care	The health state of the participant: regarding self care	I have some problems washing or dressing myself	19	0%
		I am unable to wash or dress myself	12	
		I have no problems with performing my usual activities	1161	
EQ-5D - Usual Activity	The health state of the participant: regarding activity	I have some problems with performing my usual activities	49	0%
		I am unable to perform my usual activities	2	
EQ-5D - Pain	The health state of the participant: regarding pain	I have no pain or discomfort	1027	0%

		I have moderate pain or discomfort	181	
		I have extreme pain or discomfort	4	
EQ 5D Anviotu		I am not anxious or depressed	1156	
EQ-5D - Anxiety or Depression	The health state of the participant: regarding fear and/or depression	I am moderately anxious or depressed	53	0%
of Depression		I am extremely anxious or depressed	3	
	The health state of the participant on a scale from 0 (worst) to 100 (fittest) (VAS)	Mean	87.93	
		Median	90	20/
VAS		Min	7	2%
		Max	100	
		Max	100	

^{*a*}(or is too young for self care)

Figure 1. Distribution of Average Household Health Utility Scores among Households with at least 1 member with Health Utility Score < 1 (n = 153)



At baseline, there were 257 households in which all parents were working full- or part-time with an average household health utility score of 0.95 [0.61, 1.00]. The 30 households in which at least one parent was not working full- or part-time had an average health utility score of 0.89 [0.63, 1.00]. A total of 216 households had at least one parent with at least a university degree with an average household health utility score of 0.95 [0.61, 1.00]. The 85 households in which both parents had lower than a university level education had an average health utility score of 0.92 [0.63, 1.00]. In addition to frequency counts, Table 2 outlines the results of bivariate linear models evaluating average household health utility in relation to household-level determinants. Households with one or more parents not working full or part-time [b = -0.05, p < 0.001] and households with all parents with less than a university level education [b = -0.03, p < 0.05] were significantly associated with lower average household health utility scores. Households with at least one member who has dealt with severe illness with self, family, and/or while caring for others was also found to have a significant negative effect on average household health utility scores [b = -0.04, p < 0.001].

Table 2 . Results of Bivariate Linear Regression Models Evaluating the Relationship
Between Average Household EQ-5D Scores and Household-Level Determinants

	Levels	Ν	Estimate (SE)
Household size	4 (reference)	150	0.952 (0.007) ***
	2	10	-0.037 (0.026)
	3	79	-0.039 (0.011) ***
	5	55	0.003 (0.013)
	6	11	-0.026 (0.025)
	7	2	-0.067 (0.057)
Household Parental Type	Dual-parent household (reference)	283	0.942 (0.005) ***
	Single-parent household	24	-0.024 (0.017)
Household Child Care	<i>At least one child in the household is in daycare or school (reference)</i>	301	0.875 (0.036) ***
	No children in household in daycare or school	5	0.066 (0.036) .
Household Severe Disease	No members in the household has dealt with severe illness (reference)	223	0.967 (0.009) ***
	At least one member of household has dealt	74	-0.037 (0.011) ***

with severe illness (with self, family, or because of caring for others)		
<i>There are no current smokers in the household (reference)</i>	225	0.941 (0.005) ***
At least one member of household is a current smoker	82	-0.003 (0.01)
<i>There are no pets in the household (reference)</i>	208	0.938 (0.008) ***
There is a pet in the household	99	0.003 (0.01)
Both parents working (reference)	257	0.947 (0.005) ***
At least one parent not working	30	-0.054 (0.015) ***
Both parents with less than university education	85	-0.026 (0.01) *
<i>At least one parent with university education (reference)</i>	216	0.948 (0.005) ***
	 because of caring for others) There are no current smokers in the household (reference) At least one member of household is a current smoker There are no pets in the household (reference) There is a pet in the household Both parents working (reference) At least one parent not working Both parents with less than university education At least one parent with university 	because of caring for others)There are no current smokers in the household (reference)225At least one member of household is a current smoker82There are no pets in the household (reference)208There is a pet in the household99Both parents working (reference)257At least one parent not working30Both parents with less than university education85At least one parent with university216

p < 0.001: '***'; p < 0.01: '**'; p < 0.05: '*'; p < 0.1: '.

The results of the mixed-effects models assessing the relationship between individual health utility score and the household-level determinants, while accounting for household as a clustering variable, are provided in the Supplement (<u>Suppl. Table 2</u>).

Regression Results

The final regression model is shown in <u>Table 3</u>. The pooled random effects variance for the 10 imputed datasets was 0.002 and the average ICC was 11.7% (range: 11.1%, 12.0%), indicating minimal heterogeneity in health utility scores between households and suggesting substantial homogeneity within households.

The model includes individual-level demographic determinants such as age group and occupation and the health determinant BMI. Although insignificant, children aged 12 and under appeared to have a weak negative association with health utility score (b = -0.004, p > 0.1). An unadjusted model with the binary age variable and health utility score, while accounting for household clusters, showed a weak, but positive association with health utility score for children aged 12 and under (b = 0.02, p < 0.05). The binary age classifier was selected over the

participants' numeric age due to the bimodal distribution of age in our study sample and after further comparison of unadjusted models and evaluations for multicollinearity.

Table 3 - Results of Mixed Effect Models Linking EQ-5D Scores to Individual andHousehold-Level Survey Responses

	Estimate (SE)		
Intercept	0.981 (0.008)	***	
Adult	refere	ence	
Child	-0.004 (0.033)		
Occupation : Full- or part-time employment	refere	ence	
Occupation : In full time or further education	0.002 (0.013)		
Occupation : At home (housewife), unemployed/job seeking, other	-0.066 (0.018)	***	
BMI : Normal	refere	ence	
BMI : Not Calculated	0 (0.027)		
BMI : Obesity	-0.072 (0.021)	***	
BMI : Overweight	-0.027 (0.011)	*	
BMI : Underweight	-0.037 (0.017)	*	
Exposure to Severe Disease - Caregiving : No	reference		
Exposure to Disease - Caregiving : Not applicable/Children	0.002 (0.034)		
Exposure to Disease - Caregiving : Yes	-0.026 (0.018)		
Exposure to Disease - Family : No	refere	ence	
Exposure to Disease - Family : Yes	-0.027 (0.008)	**	
Severe Disease - Self : No	refere	ence	
Severe Disease - Self : Yes	-0.062 (0.013)	***	
Uncommon Day : No	refere	ence	
Uncommon Day : Yes	-0.023 (0.008)	**	
Variance of Residuals	0.013		
Variance of Random Effect	0.0	02	
ICC	0.1	17	
Marginal R2	0.00	84	

Conditional R2	0.191
AIC	-1558.86

p < 0.001: '***'; p < 0.01: '**'; p < 0.05: '*'; p < 0.1: '.'

Participants' occupation status as 'in full-time or further education' did not show a significant effect on health utility scores (b = .002, p > 0.1); however, those with 'other' occupational status (i.e., housewife, unemployed, or job-seeking; b = -0.07, p < 0.001) were found to have a significant negative association. Participants experiencing an uncommon day at the time of survey completion (e.g. due to illness, no school or child care, or on holiday) were captured in our model and were also found to have a significant negative association was the only physical health indicator captured in our model selection. Compared to being classified as having normal BMI, underweight (b = -0.04, p < 0.05), overweight (b = -0.03, p < 0.05), and obese (b = -0.07, p < 0.001) BMI classifications were significantly associated with lower health utility scores.

When evaluated against health utility scores individually, the three survey questions regarding the participants' proximity to severe disease were shown to have negative associations with HRQoL. Being sick with a severe disease had the strongest effect on lower health utility scores (b = -0.06, p < 0.001), while having a family member with severe disease had a weaker negative effect on health utility score (b = -0.03, p < 0.01). Being a caregiver for someone with or experienced severe disease had a small, but negative association with health utility score (b = -0.04, p < 0.05) compared to not having dealt with severe illness (Suppl. Table 1). In our adjusted model, being sick with a severe disease (b = -0.06, p < 0.001) and being in close contact with a family member with a severe disease (b = -0.03, p < 0.001) remained negatively associated with lower health utility scores, while adult caregivers, regardless of exposure to severe disease, were not significantly associated with health utility scores.

Additional models were run to include household size and household smoking status, a potential indicator for health of the household. Compared to 4-member households, the median in our study sample, households with 3 members were significantly associated with lower health utility scores [b = -0.03, p < 0.01]. Having at least one smoker in the household showed a weaker, non-significant negative association with health utility score [b = -0.002, p > 0.1].

Discussion

To our knowledge, this survey is the first of its kind to collect demographic, health, and HRQoL data on households with children and to subsequently model HRQoL within and between households. Households with a higher percentage of members having health utility scores below 1 had lower average household health utility scores. On an individual level, there is minimal homogeneity in health utility scores within households. Our results indicate a non-negligible contribution of household clusters on HRQoL and individuals' occupation, physical health, and proximity to severe disease can negatively affect their HRQoL.

The primary purpose of this survey was to study social contact behaviours of household members within and outside their household. (3). A 2018 article by Goeyvaerts et al that evaluated the accompanying contact survey found a high degree of within-household clustering of contacts but, especially on weekdays, decreasing connectedness with increasing household size. (24). Since this is of interest to contact behavior, a question about whether the participants experienced a normal week or weekend day was included and, in this analysis, this was found to have a significant negative effect on EQ-5D scores. Out of the 361 respondents who reported not experiencing a normal day at baseline, 6% reported being ill, 1% reported having to stay home for a household member who was ill, and 93% cited other reasons, such as vacation or school closures. Changes in a single day's activity are not routinely ascertained in HRQoL studies. While illness has been directly associated with lower HRQoL (25), our survey data is not sufficient to draw conclusions in line with the existing literature on parents' and children's perceptions of family holidays (26,27).

Our results showed a statistically significant lower health utility score for participants at home, unemployed or job-seeking compared to the employed. At the household-level, we also found that households with at least one parent not working were significantly associated with lower average household health utility scores. These results are consistent with several existing studies (28,29), including a 2019 study by Norström et al (30), which concluded that unemployment can have a substantial impact on HRQoL. Their evaluation of unemployment against the five EQ-5D dimensions showed a significant increase in difficulties with usual activities and anxiety/depression. In our study sample, 9 participants categorized as at home or unemployed reported some problems with usual activities and 4 participants reported being somewhat anxious or depressed.

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We considered four survey questions pertaining to the participants' current health state: height, weight, smoking status, and having dealt with severe disease. The participants' BMI was calculated based on the recorded height and weight; however, for children under 2 years old, both the World Health Organization and Centre for Disease Control recommend the use of growth charts and therefore categorized as 'BMI Not Calculated'. Our model captured the participants' BMI status with all non-normal categories being significantly associated with lower health utility scores. Several studies have shown an association between obesity and HRQoL (31–33); however, Stephenson et al (34) emphasize the role of comorbidities along the causal pathway between obesity and low HRQoL.

The smoking status of adult survey participants was used to determine the household smoking status and intended to serve as a representation of the household's overall health. Smoking has been widely associated with lower HRQoL (35–37). Although not a significant factor in our model, parental smoking has been associated with childhood obesity (38), asthma and other respiratory illnesses (39) and can increase likelihood of nicotine adoption and dependence in adulthood (40).

A household member having dealt with severe illness had a significant, negative effect on average household health utility scores. When evaluated at the individual level, having a severe disease is among the factors with the strongest association with lower health utility score in our sample. However, the nature of our survey question does not delineate specific severe diseases. A 2019 systematic review by Van Wilder et al (41) compiled the EQ-5D scores related to several chronic diseases and showed that EQ-5D scores decreased in relation to increased severity of disease. The extent of severe illness can have spillover effects among family members, with effect magnitudes varying based on the closeness of familial relationships (42). The weak, but significant effect of having dealt with a family member with severe disease on lower health utility scores is suggestive of some health burden impacts in our study sample. Wu et al (43) found a positive association between the HRQoL of children with rare genetic disorders and their parents, with a calculated reduction in HRQoL of 0.06 in parents compared to parents with children without rare genetic disease. Sjolander et al. (44) evaluated the HRQoL of cancer patients' family members and showed that their HRQoL was lower than the population norm during a 1 year period after the family member's cancer diagnosis, with a pronounced effect on partners compared to children.

Exposure to severe disease as a caregiver did not have a significant effect on health utility scores. The survey question was limited to participants aged 13 and older and does not differentiate between caregiving within or beyond the household. Among the 63 respondents who reported exposure to severe disease because of caring for others, 67% worked in the public health sector. Literature on the general health related quality of life among healthcare workers is sparse (20). An Italian study on 324 healthcare workers found lower quality of life scores in vitality, social function, and emotional interference in usual activities, but higher quality of life scores in physical aspects of well-being compared to the general population (45). The authors suggest that professional caregivers are better equipped to manage the stresses of caregiving.

Similar to the question regarding participants' severe disease, the binary response to having dealt with severe illness in the family or while caring for someone else did not delineate the severity of diseases nor did it directly quantify the number of participants who are formal or informal caregivers. The heterogeneity in household-level responses to our survey question on having dealt with a family member with severe illness could also be attributed to varying interpretations of "family member" extending beyond the immediate family. Informal caregiver burden has been widely investigated (6). A 2018 survey on informal caregiving in Belgium found that 3.6% of participants provided care for patients within their own households, while 7.4% cared for family members living outside their households (17).

There are additional limitations inherent to our study. The elderly population is absent from our study sample as it was beyond the scope of the primary purpose of this survey. Buckinx et al (46) evaluated the quality of life of elderly informal caregivers and found significant mental burden and decreased physical activities compared to the general elderly population. Alternatively, Shin et al found considerable HRQoL burden for family caregivers living with elderly dementia patients compared with the quality of life of non-caregivers. Both studies highlight the significant role of elderly family members in familial health spillover effects.

In our household survey, the generic EQ-5D-3L questionnaire was given to all study participants, including children and teenagers. For children under 12 years, responses were recorded by parents, while most teenagers were asked to complete the survey on their own. In 2006, Ravens-Sieberer et al. (47) identified several challenges associated with assessing children and adolescents' HRQoL, one of which questioning the value of 'self-rated' versus 'externally rated' HRQoL measurements. A 2010 article by Wille (48) suggests that the generic EQ-5D-3L

measure was satisfactory in measuring children and adolescent HRQoL. This intrument's use, however, is limited to children at least 8 years of age. Out of our total number of study participants, 21% were younger than 8 years old and for whom Ravens-Sieberer's initial proposition is wholly relevant. Subsequent studies conducted in Germany evaluated parent-child agreement in HRQoL measurements and showed an increase in response disagreement with a decrease in child age (49). These studies also suggest external factors, such as socioeconomic status, can influence parents' estimation of their child's HRQoL (50), but are beyond the scope of our analysis.

Household-level variables were not directly ascertained in our survey; therefore, household-level determinants considered in our analysis were extrapolated from individual responses, limiting the strength of our assertions. Furthermore, the broad categorization of our occupation variable grouped participants working full- or part-time, overlooking the potential income-related nuances that may affect HRQoL at the individual and household-level. Education is often used alongside occupation as an indicator for socioeconomic status (51). Variations of individual education levels were considered for this analysis and notably excluded from the final model. At the household-level, we found that households in which both parents did not have a university degree were significantly associated with lower household average health utility scores. In Belgium, both salaries and health utility scores have been shown to increase with higher educational attainment (20,52,53). Additionally, Van Droogenbroek et al. (17) found significant associations between household income and emotional distress as a result of providing informal care.

Household-level demographic factors and exposure to severe disease are significantly related to average household health utility. The moderate impact of household clusters in our model's results additionally suggests a non-negligible effect on individual health utility that requires further investigation. It is important to note that individual level circumstances only partially captured by our survey, such as provision of care to external family members, may still lead to variability within households. Our survey captured participants' health utility scores in a single time point in their lifetime, but evidence suggests that prolonged positive or negative health effects can have downstream implications to an individual's close social relationships.

References

- 1. World Health Organization. Constitution of the World Health Organization [Internet]. 2024 [cited 2024 Oct 21]. Available from: https://www.who.int/about/governance/constitution
- Holt-Lunstad J, Smith TB, Layton JB. Social Relationships and Mortality Risk: A Meta-analytic Review. PLoS Med. 2010 Jul 27;7(7):e1000316.
- 3. Goeyvaerts N, Santermans E, Potter G, Torneri A, Van Kerckhove K, Willem L, et al. Household members do not contact each other at random: implications for infectious disease modelling. Proc R Soc B Biol Sci. 2018 Dec 19;285(1893):20182201.
- Hoang TV, Coletti P, Kiffe YW, Kerckhove KV, Vercruysse S, Willem L, et al. Close contact infection dynamics over time: insights from a second large-scale social contact survey in Flanders, Belgium, in 2010-2011 [Internet]. medRxiv; 2020 [cited 2023 Apr 24]. p. 2020.09.30.20204891. Available from:

https://www.medrxiv.org/content/10.1101/2020.09.30.20204891v1

- 5. Prosser LA, Lamarand K, Gebremariam A, Wittenberg E. Measuring Family HRQoL Spillover Effects Using Direct Health Utility Assessment. Med Decis Making. 2015 Jan 1;35(1):81–93.
- Wittenberg E, James LP, Prosser LA. Spillover Effects on Caregivers' and Family Members' Utility: A Systematic Review of the Literature. PharmacoEconomics. 2019 Apr 1;37(4):475–99.
- 7. Henry E, Cullinan J. Mental health spillovers from serious family illness: Doubly robust estimation using EQ-5D-5L population normative data. Soc Sci Med. 2021 Jun 1;279:113996.
- 8. Wittenberg E, Prosser LA. Disutility of illness for caregivers and families: a systematic review of the literature. PharmacoEconomics. 2013 Jun;31(6):489–500.
- 9. Wittenberg E, Ritter G, Prosser LA. Evidence of spillover effects of illness among household members: EQ-5D scores from a US population sample. Med Decis Mak Int J Soc Med Decis Mak. 2013 Feb;33(2):235–43.
- 10. Wittenberg E, Saada A, Prosser LA. How illness affects family members: a qualitative interview survey. The patient. 2013 Dec;6(4):10.1007/s40271-013-0030–3.
- 11. Theofilou P, Vlastos DD. The Psychological Burden of Families with Diabetic Children: A Literature Review Focusing on Quality of Life and Stress. Children. 2023 Jun;10(6):937.
- 12. Haugstvedt A, Wentzel-Larsen T, Rokne B, Graue M. Perceived family burden and emotional distress: similarities and differences between mothers and fathers of children with type 1 diabetes in a population-based study. Pediatr Diabetes. 2011 Mar;12(2):107–14.
- 13. Molla IB, Berhie MA, Debele KA, Germossa GN, Hailu FB. Persons with Diabetes' Perceptions of Family Burden and Associated Factors. J Diabetes Res. 2023 Jan 5;2023:8015721.
- Radicke A, Barkmann C, Adema B, Daubmann A, Wegscheider K, Wiegand-Grefe S. Children of Parents with a Mental Illness: Predictors of Health-Related Quality of Life and Determinants of Child-Parent Agreement. Int J Environ Res Public Health. 2021 Jan 6;18(2):379.
- 15. Lawson KM, Davis KD, McHale SM, Hammer LB, Buxton OM. Daily Positive Spillover and Crossover from Mothers' Work to Youth Health. J Fam Psychol JFP J Div Fam Psychol Am Psychol Assoc Div 43. 2014 Dec;28(6):897–907.
- 16. Gerkens S, Lefèvre M, Bouckaert N, Levy M, Noordhout CM de, Obyn C, et al. Performance of the Belgian health system: Report 2024. Brussels: Belgian Health Care

Knowledge Centre (KCE); 2024. (KCE Reports). Report No.: 376C.

- 17. Van Droogenbroeck F, Spruyt B, Gérain P, Van den Borre L, Smith P, De Pauw R, et al. Informal caregiving and mental health: results from the Belgian health interview survey 2013 and 2018. BMC Public Health. 2025 Jan 2;25(1):15.
- 18. Saloni Dattani, Lucas Rodés-Guirao, Hannah Ritchie, Esteban Ortiz-Ospina. OurWorldinData.org. 2023 [cited 2025 Jan 31]. Life Expectancy. Available from: https://ourworldindata.org/life-expectancy
- 19. Nakanishi M, Stanyon D, Richards M, Yamasaki S, Ando S, Endo K, et al. Informal Caregiving in Adolescents from 10 to 16 Years Old: A Longitudinal Study Using Data from the Tokyo Teen Cohort. Int J Environ Res Public Health. 2023 Jul 31;20(15):6482.
- 20. Bilcke J, Hens N, Beutels P. Quality-of-life: a many-splendored thing? Belgian population norms and 34 potential determinants explored by beta regression. Qual Life Res. 2017 Aug 1;26(8):2011–23.
- 21. Cleemput I. A social preference valuations set for EQ-5D health states in Flanders, Belgium. Eur J Health Econ. 2010 Apr 1;11(2):205–13.
- 22. van Buuren S. Flexible Imputation of Missing Data, Second Edition. 2nd ed. New York: Chapman and Hall/CRC; 2018. 444 p.
- 23. Molenberghs G, Verbeke G. Models for Discrete Longitudinal Data [Internet]. 1st ed. New York: Springer-Verlag; 2005 [cited 2024 Oct 21]. XXII, 687. (Springer Series in Statistics). Available from: http://link.springer.com/10.1007/0-387-28980-1
- 24. Nesti MM, Goldbaum M. Infectious diseases and daycare and preschool education. J Pediatr (Rio J). 2007;83(4):299–312.
- 25. Devlin N, Parkin D, Janssen B. Methods for Analysing and Reporting EQ-5D Data [Internet]. Cham (CH): Springer; 2020 [cited 2025 Feb 10]. Available from: http://www.ncbi.nlm.nih.gov/books/NBK565678/
- 26. Fu X, Lehto X, Park O. What Does Vacation do to our Family? Contrasting the Perspectives of Parents and Children. J Travel Tour Mark. 2014 May 19;31(4):461–75.
- 27. Backer E, Schänzel H. Family Holidays—Vacation or Obli-cation? Tour Recreat Res. 2013 Jan 1;38(2):159–73.
- 28. Feng YS, Kohlmann T, Janssen MF, Buchholz I. Psychometric properties of the EQ-5D-5L: a systematic review of the literature. Qual Life Res Int J Qual Life Asp Treat Care Rehabil. 2021 Mar;30(3):647–73.
- 29. Gautier L, Azzi J, Saba G, Bonnelye G, de Pouvourville G. Population norms in France with EQ-5D-5L: health states, value indexes, and VAS. Eur J Health Econ. 2023 Dec 1;24(9):1517–30.
- 30. Norström F, Waenerlund AK, Lindholm L, Nygren R, Sahlén KG, Brydsten A. Does unemployment contribute to poorer health-related quality of life among Swedish adults? BMC Public Health. 2019 Apr 29;19(1):457.
- 31. Schwimmer JB, Burwinkle TM, Varni JW. Health-Related Quality of Life of Severely Obese Children and Adolescents. JAMA. 2003 Apr 9;289(14):1813–9.
- 32. Mejaddam A, Krantz E, Höskuldsdóttir G, Fändriks L, Mossberg K, Eliasson B, et al. Comorbidity and quality of life in obesity–a comparative study with the general population in Gothenburg, Sweden. PLOS ONE. 2022 Oct 4;17(10):e0273553.
- 33. Busutil R, Espallardo O, Torres A, Martínez-Galdeano L, Zozaya N, Hidalgo-Vega Á. The impact of obesity on health-related quality of life in Spain. Health Qual Life Outcomes. 2017 Oct 10;15(1):197.

- 34. Stephenson J, Smith CM, Kearns B, Haywood A, Bissell P. The association between obesity and quality of life: a retrospective analysis of a large-scale population-based cohort study. BMC Public Health. 2021 Nov 3;21(1):1990.
- 35. Cheng X, Jin C. The Association Between Smoking and Health-Related Quality of Life Among Chinese Individuals Aged 40 Years and Older: A Cross-Sectional Study. Front Public Health. 2022 Feb 24;10:779789.
- 36. Mitra M, Chung MC, Wilber N, Klein Walker D. Smoking status and quality of life: a longitudinal study among adults with disabilities. Am J Prev Med. 2004 Oct;27(3):258–60.
- 37. Cui Y, Forget EL, Torabi M, Oguzoglu U, Ohinmaa A, Zhu Y. Health-related quality of life and economic burden to smoking behaviour among Canadians. Can J Public Health Rev Can Sante Publique. 2019 Oct;110(5):533–41.
- 38. Srivastava P, Trinh TA, Hallam KT, Karimi L, Hollingsworth B. The links between parental smoking and childhood obesity: data of the longitudinal study of Australian children. BMC Public Health. 2024 Jan 2;24(1):68.
- 39. Pattenden S, Antova T, Neuberger M, Nikiforov B, De Sario M, Grize L, et al. Parental smoking and children's respiratory health: independent effects of prenatal and postnatal exposure. Tob Control. 2006 Aug;15(4):294–301.
- 40. Mays D, Gilman SE, Rende R, Luta G, Tercyak KP, Niaura RS. Parental Smoking Exposure and Adolescent Smoking Trajectories. Pediatrics. 2014 Jun;133(6):983–91.
- 41. Van Wilder L, Rammant E, Clays E, Devleesschauwer B, Pauwels N, De Smedt D. A comprehensive catalogue of EQ-5D scores in chronic disease: results of a systematic review. Qual Life Res. 2019 Dec 1;28(12):3153–61.
- 42. Lavelle TA, D'Cruz BN, Mohit B, Ungar WJ, Prosser LA, Tsiplova K, et al. Family Spillover Effects in Pediatric Cost-Utility Analyses. Appl Health Econ Health Policy. 2019 Apr 1;17(2):163–74.
- 43. Wu Y, Al-Janabi H, Mallett A, Quinlan C, Scheffer IE, Howell KB, et al. Parental health spillover effects of paediatric rare genetic conditions. Qual Life Res. 2020 Sep 1;29(9):2445–54.
- 44. Sjolander C, Rolander B, Järhult J, Mårtensson J, Ahlstrom G. Health-related quality of life in family members of patients with an advanced cancer diagnosis: A one-year prospective study. Health Qual Life Outcomes. 2012 Jul 30;10:89.
- 45. Kheiraoui F, Gualano MR, Mannocci A, Boccia A, La Torre G. Quality of life among healthcare workers: A multicentre cross-sectional study in Italy. Public Health. 2012 Jul 1;126(7):624–9.
- 46. Buckinx F, Adam S, Aubertin-Leheudre M, De Saint Hubert M, Mouton A, Potier F, et al. Quality of Life and Health Determinants of Informal Caregivers Aged 65 Years and Over. Epidemiologia. 2023 Nov 6;4(4):464–82.
- 47. Ravens-Sieberer U, Erhart M, Wille N, Wetzel R, Nickel J, Bullinger M. Generic health-related quality-of-life assessment in children and adolescents: methodological considerations. PharmacoEconomics. 2006;24(12):1199–220.
- 48. Wille N, Badia X, Bonsel G, Burström K, Cavrini G, Devlin N, et al. Development of the EQ-5D-Y: a child-friendly version of the EQ-5D. Qual Life Res. 2010;19(6):875–86.
- 49. Erhart M, Ellert U, Kurth BM, Ravens-Sieberer U. Measuring adolescents' HRQoL via self reports and parent proxy reports: an evaluation of the psychometric properties of both versions of the KINDL-R instrument. Health Qual Life Outcomes. 2009 Aug 26;7:77.
- 50. Ellert U, Ravens-Sieberer U, Erhart M, Kurth BM. Determinants of agreement between

self-reported and parent-assessed quality of life for children in Germany-results of the German Health Interview and Examination Survey for Children and Adolescents (KiGGS). Health Qual Life Outcomes. 2011 Nov 23;9(1):102.

- 51. CDC. Centers for Disease Control and Prevention. 2023 [cited 2025 Jan 27]. Socioeconomic Factors | CDC. Available from: https://www.cdc.gov/dhdsp/health equity/socioeconomic.htm
- 52. STATBEL. Directorate-general Statistics Statistics Belgium. 2024 [cited 2025 Jan 27]. Level of education. Available from: https://statbel.fgov.be/en/themes/census/education/level-education
- 53. Sciensano. Life Expectancy and Quality of Life: Quality of life [Internet]. Brussels, Belgium; 2022 Jan [cited 2025 Jan 29]. Available from: https://www.healthybelgium.be/en/health-status/life-expectancy-and-quality-of-life/quality-of-life

Appendix

Regression Results

Supplementary Table 1 - Pooled Results from the Linear Mixed-Effects Models Evaluating the Relationship between Individual Health Utility Score and the Individual-Level Determinants, while Accounting for Household as a Clustering Variable

				Variance of	
	Variable	Estimate (S	SE)	Random Effect	ICC
	Adult (reference)	0.93 (0.01)	***	0.0024	0.1465
Туре	Teenager or older child himself	0.03 (0.01)	*		
	Parent on behalf of child	0.02 (0.01)	**		
A go	(Intercept)	0.96 (0.01)	***	0.0023	0.1438
Age	Participant age	-0.001 (0)	**		
Binary Age	Adult (reference)	0.935 (0.01)	***	0.0023	0.1419
Group Category	Child	0.016 (0.01)	*		
Gender	Male (reference)	0.941 (0.01)	***	0.0023	0.1431
	Female	0.002 (0.01)			
Uncommon	No (reference)	0.95 (0.01)	***	0.0023	0.1383
Day	Yes	-0.02 (0.01)	*		
Child School	Child in primary school (reference)	0.969 (0.01)	***	0.0022	0.1390
	Adult or no response	-0.034 (0.01)	***		
	Child in preschool	-0.033 (0.01)	*		
	Child in school; not pre- or primary school	-0.037 (0.02)	*		
	Child not in school	-0.054 (0.02)	**		
Child School	Child in school (reference)	0.96 (0.01)	***	0.0023	0.1413
(Recategorized)	Adult or no response	-0.02 (0.01)	**		
(Recategorized)	Child not in school	-0.04 (0.02)	*		
Child Care	Child attending school or daycare (reference)	0.95 (0.01)	***	0.0023	0.1389
	Adult or no response	-0.02 (0.01)	*		
	Child not attending school or daycare	-0.04 (0.04)			
Education	University degree (lower) (reference)	0.92 (0.01)	***	0.0023	0.1400
	No formal schooling	0.01 (0.06)			
	Primary school	0.02 (0.02)			
	Secondary school (lower)	0.02 (0.02)			

	Upper secondary school (upper)	0.02 (0.02)			
	Secondary school (unspecified)	0.03 (0.02)			
	University degree (higher)	0.03 (0.02)			
Occupation	Working full- or part-time (reference)	0.94 (0.01)	***	0.0023	0.1407
	At home (housewife)	-0.03 (0.03)			
	Currently unemployed / job seeking	-0.11 (0.03)	**		
	In full time or further education	0.02 (0.01)	**		
	Other	-0.03 (0.02)			
Severe Disease	Did not deal with a serious illness (reference)	0.95 (0)	***	0.0021	0.1303
- Self	Have dealt with a serious illness	-0.07 (0.01)	***		
Severe Disease	Did not deal with a serious illness in family (reference)	0.96 (0.01)	***	0.0021	0.1290
- Family	Have dealt with a serious illness in family	-0.04 (0.01)	***		
	<i>Did not deal with a serious illness because of caring for others (reference)</i>	0.94 (0.01)	***	0.0023	0.1414
Severe Disease - Caregiving	Child; not applicable	0.01 (0.01)			
Curogiving	Have dealt with a serious illness because of caring for others	-0.04 (0.02)	*		
	Current smoker (reference)	0.91 (0.01)	***	0.0024	0.1451
	Child or no response	0.04 (0.01)	**		
Smoking Status	Ex-smoker	0.01 (0.02)			
	Non-smoker	0.03 (0.01)	*		
	Unknown	-0.06 (0.06)			
VAS	(Intercept)	0.56 (0.03)	***	0.0017	0.1216
	Constant	0.004 (0)	***		
M adama 1	University Degree (reference)	0.95 (0.01)	***	0.0023	0.1389
Maternal Education	Adult or no response	-0.02 (0.01)	*		
Education	Less than university	-0.01 (0.01)			
A 114	University Degree (reference)	0.93 (0.01)	***	0.0023	0.1410
Adult Education	Child or no response	0.02 (0.01)			
Education	Less than university	0 (0.01)			
	Working full- or part-time (reference)	0.94 (0.01)	***	0.0023	0.1443
Occupation -	In full time or further education	0.02 (0.01)	**		
Recategorized	Stay at-home parent, unemployed, job-seeking, or other	-0.04 (0.01)	**		

Education - Recategorized of	University (reference)	0.94 (0.01)	***	0.0024	0.1467
	Unknown	-0.04 (0.07)			
	Child, not in school/child in preschool or daycare	-0.01 (0.01)			
	^a Primary school	0.03 (0.01)	**		
	Secondary school	0 (0.01)			
	No formal schooling	0.03 (0.07)			
Height	(Intercept)	0.96 (0.02)	***	0.0024	0.1433
	Height	-0.01 (0.01)			
Weight	(Intercept)	0.96 (0.01)	***	0.0023	0.1427
	Weight	-0.0003 (0.0001)) *		
BMI - Continuous	(Intercept)	1 (0.02)	***	0.0023	0.1429
	BMI	-0.003 (0.001)	***		
BMI - Categorical	Normal (reference)	0.95 (0)	***	0.0023	0.1454
	Not Calculated/Infants	-0.05 (0.02)	*		
	Underweight	-0.03 (0.02)			
	Overweight	-0.04 (0.01)	***		
	Obesity	-0.08 (0.02)	***		

p < 0.001: '***'; p < 0.01: '**'; p < 0.05: '*'; p < 0.1: '.'

^aTransformed education combines child and adult education responses based on 'child school', 'education (original)', 'response type', and 'age group'.

Supplementary Table 2 - Pooled Results from the Linear Mixed-Effects Models Evaluating the Relationship between Individual Health Utility Score and the Household-Level Determinants, while Accounting for Household as a Clustering Variable

	Levels	Estimate (SE)	Variance of Random Effect	ICC
Household size	4 (reference)	0.952 (0.006) ***	0.0021	0.1321
	2	-0.037 (0.031)		
	3	-0.039 (0.011) ***		
	5	0.003 (0.011)		
	6	-0.026 (0.021)		
	7	-0.067 (0.046)		
Household Parental Type	Dual-parent household (reference)	0.94 (0) ***	0.0023	0.1425

	Single-parent family	-0.03 (0.02)		
Household Child Care	At least one child in the household is in daycare or school (reference)	0.942 (0.004) ***	0.0024	0.1450
	No	-0.046 (0.037)		
Household Severe Disease	No members in the household has dealt with severe illness (reference)	0.968 (0.009) ***	0.0021	0.1323
Disease	Yes	-0.035 (0.01) ***		
Household Smoking	<i>There are no current smokers in the household (reference)</i>	0.942 (0.005) ***	0.0024	0.1450
	Yes	-0.002 (0.01)		
Household Animal Ownership	<i>There are no pets in the household (reference)</i>	0.934 (0.017) ***	0.0024	0.1445
	Yes	0.004 (0.01)		

p < 0.001: '***'; p < 0.01: '**'; p < 0.05: '*'; p < 0.1: '.'

Compared to the average household size of 4, households with 3 members were significantly associated with lower health utility scores [b = -0.05, p < 0.001]. Households with at least one member having dealt with a serious illness (with self, family, and/or while caring for others) was significantly associated with lower health utility scores [b = -0.04, p < 0.001]. Single-parent households, households without access to child care and with at least one member currently smoking had lower, but not significant, health utility scores, while having an animal in the household was associated with a minimal, but not significant, increase in health utility scores.